

## Introduction to Brownian Motion

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### Abstract

This paper presents the basic knowledge of a standard Brownian motion which is a building block of all stochastic processes. A standard Brownian motion is a subclass of 1) continuous martingales, 2) Markov processes, 3) Gaussian processes, and 4) Itô diffusion processes. It is also a subclass of Lévy processes although we will not discuss this in this sequel.

## [1] Brownian Motion

### [1.1] Standard Brownian Motion

**Definition 1.1 Standard Brownian motion (Standard Wiener process)** A standard Brownian motion  $(B_{t \in [0, \infty)})$  is a real valued stochastic process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  satisfying:

(1) Its increments are independent. In other words, for  $0 \leq t_1 < t_2 < \dots < t_n < \infty$ :

$$\begin{aligned} & \mathbb{P}(B_{t_0} \cap B_{t_1} - B_{t_0} \cap B_{t_2} - B_{t_1} \cap \dots \cap B_{t_n} - B_{t_{n-1}}) \\ &= \mathbb{P}(B_{t_0}) \mathbb{P}(B_{t_1} - B_{t_0}) \mathbb{P}(B_{t_2} - B_{t_1}) \dots \mathbb{P}(B_{t_n} - B_{t_{n-1}}). \end{aligned}$$

(2) Its increments are stationary (time homogeneous): i.e. for  $h \geq 0$ ,  $B_{t+h} - B_t$  has the same distribution as  $B_h$ . In other words, the distribution of increments does not depend on  $t$ .

(3)  $\mathbb{P}(B_0 = 0) = 1$ . The process starts from 0 almost surely (with probability 1).

(4)  $B_t \sim \text{Normal}(0, t)$ . Its increments follow a Gaussian distribution with the mean 0 and the variance  $t$ .

**Definition 1.2 Standard Brownian motion with starting point  $c$**  Let  $c$  be a real valued constant or a random variable independent of a standard Brownian motion  $(B_{t \in [0, \infty)})$ . Then, a standard Brownian motion with starting point  $c$  is a real valued stochastic process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ :

$$(c + B_{t \in [0, \infty)}).$$

**Theorem 1.1 Standard Brownian motion** A standard Brownian motion process  $(B_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  satisfies the following conditions:

(1) The process is stochastically continuous:  $\forall \varepsilon > 0, \lim_{h \rightarrow 0} \mathbb{P}(|X_{t+h} - X_t| \geq \varepsilon) = 0$ .

(2) Its sample path (trajectory) is continuous in  $t$  (i.e. continuous  $\in$  rcll) almost surely.

Proof

Consult Karlin (1975). We have to remind you that this proof is not that simple.

### [1.2] Brownian Motion with Drift

**Definition 1.3 Brownian motion with drift** Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, a Brownian motion with drift is a real valued stochastic process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  as:

$$(X_{t \in [0, \infty)}) \equiv (\mu t + \sigma B_{t \in [0, \infty)}),$$

where  $\mu \in \mathbb{R}$  is called a drift and  $\sigma \in \mathbb{R}^+$  is called a diffusion (volatility) parameter. A Brownian motion with drift satisfies the following conditions:

(1) Its increments are independent. In other words, for  $0 \leq t_1 < t_2 < \dots < t_n < \infty$ :

$$\begin{aligned} & \mathbb{P}(X_{t_0} \cap X_{t_1} - X_{t_0} \cap X_{t_2} - X_{t_1} \cap \dots \cap X_{t_n} - X_{t_{n-1}}) \\ &= \mathbb{P}(X_{t_0}) \mathbb{P}(X_{t_1} - X_{t_0}) \mathbb{P}(X_{t_2} - X_{t_1}) \dots \mathbb{P}(X_{t_n} - X_{t_{n-1}}). \end{aligned}$$

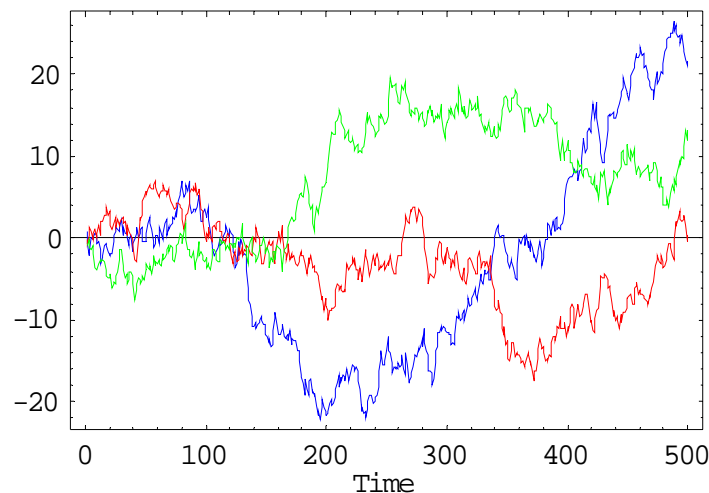
(2) Its increments are stationary (time homogeneous): i.e. for  $h \geq 0$ ,  $X_{t+h} - X_t$  has the same distribution as  $X_h$ . In other words, the distribution of increments does not depend on  $t$ .

(3)  $X_t \equiv \mu t + \sigma B_t \sim Normal(\mu t, \sigma^2 t)$ . Its increments follow a Gaussian distribution with the mean  $\mu t$  and the variance  $\sigma^2 t$ .

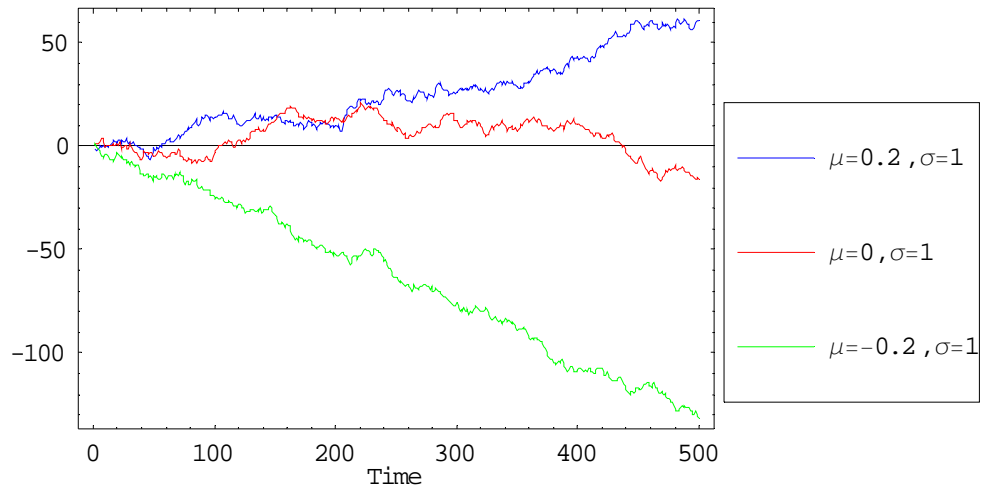
(4) Its sample path (trajectory) is continuous in  $t$  (i.e. continuous  $\in$  rcll) almost surely.

### [1.3] Sample Paths Properties of Brownian Motion

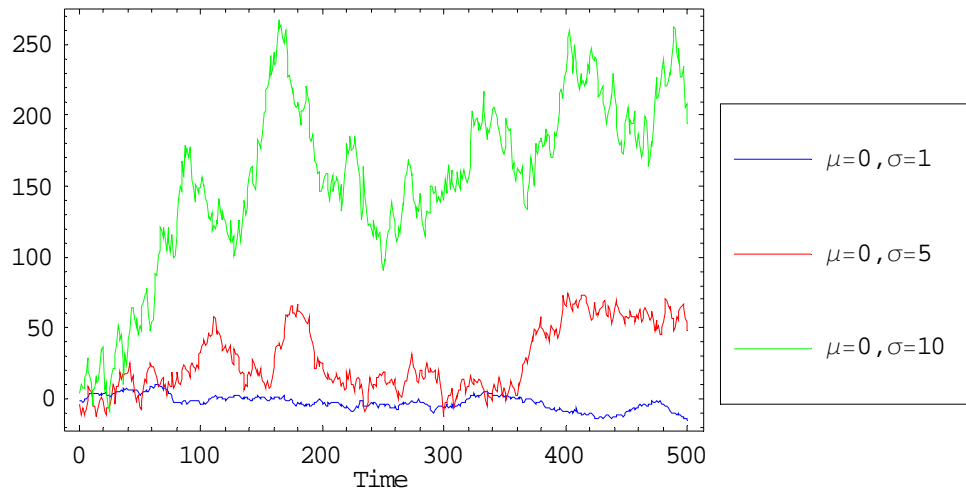
Before discussing the sample paths properties of Brownian motion, take a look at simulated sample paths of a standard Brownian motion on Panel (A) in Figure 1.1 and those of a Brownian motion with drift on Panel (B) and (C).



A) Sample Paths of a Standard Brownian Motion.



B) Sample Paths of a Brownian Motion with Drift. Different drifts and same diffusion parameters.



C) Sample Paths of a Brownian Motion with Drift. Zero drifts and different diffusion parameters.

**Figure 1.1 Simulated Sample Paths of Brownian Motion**

**Theorem 1.2 Sample paths properties of Brownian motion with drift** Consider a real valued Brownian motion with drift  $(X_{t \in [0, T]}) \equiv (\mu t + \sigma B_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, the sample paths of  $(X_{t \in [0, T]})$  possess following properties:

- (1) Sample paths are continuous with probability 1.

(2) Sample paths are of infinite variation on any finite interval  $[0, t]$ . In other words, the total variation on any finite interval  $[0, t]$  of a sample path of a Brownian motion with drift is infinite with probability 1 in the limit  $n \rightarrow \infty$  (as the partition becomes finer and finer):

$$\mathbb{P} \left( \lim_{n \rightarrow \infty} T(X) = \limsup_{n \rightarrow \infty} \sum_{i=1}^n |X(t_i) - X(t_{i-1})| = \infty \right) = 1.$$

Intuitively speaking, the infinite variation property means highly oscillatory sample paths.

(3) The quadratic variations of sample paths of Brownian motions with drift  $(X_{t \in [0, T]})$  are finite on any finite interval  $[0, t]$  and converge to  $\sigma^2 t$  with probability 1 in the limit  $n \rightarrow \infty$  (as the partition becomes finer and finer):

$$\mathbb{P} \left( \lim_{n \rightarrow \infty} T^2(X) = \limsup_{n \rightarrow \infty} \sum_{i=1}^n |X(t_i) - X(t_{i-1})|^2 = \sigma^2 t < \infty \right) = 1.$$

For more details and proofs about theorem 1.2, consult Sato (1999) page 22 – 28 and Karatzas and Shreve (1991) section 1.5 and 2.9. We also recommend Rogers and Williams (2000) chapter 1.

#### [1.4] Equivalent Transformations of Brownian Motion

**Theorem 1.3 Equivalent transformations of Brownian motion** If  $(B_{t \in [0, \infty)})$  is a real valued standard Brownian motion defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ , then, it satisfies the four conditions:

(1) A standard Brownian motion  $(B_{t \in [0, \infty)})$  is symmetric. In other words, the process  $(-B_{t \in [0, \infty)})$  is also a standard Brownian motion:

$$(B_{t \in [0, \infty)}) \stackrel{d}{=} (-B_{t \in [0, \infty)}).$$

(2) A standard Brownian motion  $(B_{t \in [0, \infty)})$  has a time shifting property. In other words, the process  $(B_{t+A} - B_A)$  is also a standard Brownian motion for  $\forall A \in \mathbb{R}^+$ :

$$(B_{t+A} - B_A) \stackrel{d}{=} (B_{t \in [0, \infty)}).$$

(3) Time scaling property of a standard Brownian motion. For any nonzero  $c \in \mathbb{R}$ , the process  $(\sqrt{c}B_{t/c})$  or  $(\frac{1}{\sqrt{c}}B_{ct})$  is also a standard Brownian motion:

$$\left(\frac{1}{\sqrt{c}} B_{ct}\right) \stackrel{d}{=} (\sqrt{c} B_{t/c}) \stackrel{d}{=} (B_{t \in [0, \infty)}).$$

(4) Time inversion property of a standard Brownian motion (i.e. a variant of (3)). The process defined as:

$$(\tilde{B}_{t \in [0, \infty)}) = \begin{cases} 0 & \text{if } t = 0 \\ (tB_{1/t}) & \text{if } 0 < t < \infty \end{cases},$$

is also a standard Brownian motion:

$$(\tilde{B}_{t \in [0, \infty)}) \stackrel{d}{=} (B_{t \in [0, \infty)}).$$

Proof

These are easy exercises for readers. For the proof of the continuity of  $(\tilde{B}_{t \in [0, \infty)})$  at 0, consult Rogers and Williams (2000) page 4.

### [1.5] Characteristic Function of Brownian Motion

Consider a real valued Brownian motion with drift process  $(X_{t \in [0, \infty)}) \equiv (\mu t + \sigma B_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Its characteristic function can be obtained by the direct use of the definition of a characteristic function (i.e. Fourier transform of the probability density function with Fourier transform parameters (1,1)):

$$\begin{aligned} \phi_{X_t}(\omega) &\equiv \mathcal{F}[\mathbb{P}(x)] \equiv \int_{-\infty}^{\infty} e^{i\omega x} \mathbb{P}(x) dx \\ \phi_{X_t}(\omega) &= \int_{-\infty}^{\infty} e^{i\omega x} \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{(x - \mu t)^2}{2\sigma^2 t}\right\} dx \\ \phi_{X_t}(\omega) &= \exp(i\mu t \omega - \frac{\sigma^2 t \omega^2}{2}). \end{aligned}$$

### [2] Brownian Motion as a Subclass of Continuous Martingale

**Definition 2.1 Continuous martingale** A continuous stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  is said to be a continuous martingale with respect to the filtration  $\mathcal{F}_t$  and under the probability measure  $\mathbb{P}$  if it satisfies the following conditions:

- (1)  $X_t$  is nonanticipating.

(2)  $E[|X_t|] < \infty$  for  $\forall t \in [0, T]$ . Finite mean condition.

(3)  $E[X_u | \mathcal{F}_t] = X_t$  for  $\forall u > t$ .

In other words, if a stochastic process is a martingale, then, the best prediction of its future value is its present value. Note that the definition of martingale makes sense only when the underlying probability measure  $P$  and the filtration  $\mathcal{F}_t$  have been specified.

The fundamental property of a martingale process is that its future variations are completely unpredictable with the filtration  $\mathcal{F}_t$ :

$$\forall u > 0, E[x_{t+u} - x_t | \mathcal{F}_t] = E[x_{t+u} | \mathcal{F}_t] - E[x_t | \mathcal{F}_t] = x_t - x_t = 0.$$

Finite mean condition (2) is necessary to ensure the existence of the conditional expectation.

## [2.1] A Continuous Martingale Property of Standard Brownian Motion

**Theorem 2.1 Standard Brownian motion is a continuous martingale** Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then,  $(B_{t \in [0, \infty)})$  is a continuous martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

Proof

By definition,  $(B_{t \in [0, \infty)})$  is a nonanticipating process (i.e.  $\mathcal{F}_{t \in [0, \infty)}$ -adapted process) with the finite mean  $E[|B_t|] = 0 < \infty$  for  $\forall t \in [0, \infty)$ . For  $\forall 0 \leq t \leq u < \infty$ :

$$B_u = B_t + \int_t^u dB_v. \quad (1)$$

Using the equation (1) and the fact that a Brownian motion is a nonanticipating process, i.e.  $E[B_t | \mathcal{F}_t] = B_t$ :

$$E[B_u - B_t | \mathcal{F}_t] = E[B_u | \mathcal{F}_t] - E[B_t | \mathcal{F}_t] = E[B_t + \int_t^u dB_v | \mathcal{F}_t] - B_t$$

$$E[B_u - B_t | \mathcal{F}_t] = E[B_t | \mathcal{F}_t] + E[\int_t^u dB_v | \mathcal{F}_t] - B_t$$

$$E[B_u - B_t | \mathcal{F}_t] = B_t + 0 - B_t = 0,$$

or in other words:

$$E[B_u | \mathcal{F}_t] = E[B_t + \int_t^u dB_v | \mathcal{F}_t] = E[B_t | \mathcal{F}_t] + E[\int_t^u dB_v | \mathcal{F}_t] = B_t + 0$$

$$E[B_u | \mathcal{F}_t] = B_t,$$

which is a martingale condition. □

**Theorem 2.2 Squared standard Brownian motion  $B_t^2$  is not a continuous martingale**

Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then,  $B_t^2$  is not a continuous martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

Proof

Using the equation (1) and independent increments condition, for  $0 \leq t \leq u < \infty$  :

$$E[B_u^2 | \mathcal{F}_t] = E[(B_t + \int_t^u dB_v)^2 | \mathcal{F}_t]$$

$$E[B_u^2 | \mathcal{F}_t] = E[B_t^2 + 2B_t \int_t^u dB_v + \int_t^u dB_v^2 | \mathcal{F}_t]$$

$$E[B_u^2 | \mathcal{F}_t] = E[B_t^2 | \mathcal{F}_t] + E[2B_t \int_t^u dB_v | \mathcal{F}_t] + E[\int_t^u dB_v^2 | \mathcal{F}_t]$$

$$E[B_u^2 | \mathcal{F}_t] = E[B_t^2 | \mathcal{F}_t] + E[\int_t^u dB_v^2 | \mathcal{F}_t]$$

$$E[B_u^2 | \mathcal{F}_t] = B_t^2 + (u - t),$$

which violates a martingale condition. □

**Theorem 2.3 Squared standard Brownian motion minus time  $B_t^2 - t$  is a continuous martingale**

Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then,  $B_t^2 - t$  is a continuous martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

Proof

Using the equation (1) and independent increments condition, for  $0 \leq t \leq u < \infty$  :

$$E[B_u^2 - u | \mathcal{F}_t] = E[(B_t^2 - t) + \{(\int_t^u dB_v)^2 - (u - t)\} | \mathcal{F}_t]$$



$$\begin{aligned}
 E[B_u^2 - u | \mathcal{F}_t] &= E[(B_t^2 - t) | \mathcal{F}_t] + E[(\int_t^u dB_v)^2 - (u - t) | \mathcal{F}_t] \\
 E[B_u^2 - u | \mathcal{F}_t] &= B_t^2 - t + E[(\int_t^u dB_v)^2 | \mathcal{F}_t] - (u - t) \\
 E[B_u^2 - u | \mathcal{F}_t] &= B_t^2 - t + (u - t) - (u - t) = B_t^2 - t,
 \end{aligned}$$

which satisfies a martingale condition. □

**Theorem 2.4 Converse of theorem 2.3** Let  $(X_{t \in [0, \infty)})$  be a continuous martingale defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, the process  $(X_{t \in [0, \infty)})$  is a standard Brownian motion if and only if it satisfies:

- (1)  $X_0 = 0$ . The process starts from zero.
- (2)  $X_t^2 - t$  is a martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

## [2.2] Nonmartingale Property of a Brownian Motion with Drift

**Theorem 2.5 Brownian motion with drift is not a continuous martingale** Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, a Brownian motion with drift  $(X_{t \in [0, \infty)}) \equiv (\mu t + \sigma B_{t \in [0, \infty)})$  is not a continuous martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

Proof

By definition,  $(X_{t \in [0, \infty)})$  is a nonanticipating process (i.e.  $\mathcal{F}_{t \in [0, \infty)}$ -adapted process) with the finite mean  $E[X_t] = E[\mu t + \sigma B_t] = \mu t < \infty$  for  $\forall t \in [0, \infty)$  and  $\mu \in \mathbb{R}$ . For  $\forall 0 \leq t \leq u < \infty$ :

$$X_u = X_t + \int_t^u dX_v. \quad (2)$$

Using the equation (2) and the fact that a Brownian motion with drift is a nonanticipating process, i.e.  $E[X_t | \mathcal{F}_t] = X_t$ :

$$\begin{aligned}
 E[X_u | \mathcal{F}_t] &= E[X_t + \int_t^u dX_v | \mathcal{F}_t] = E[X_t | \mathcal{F}_t] + E[\int_t^u dX_v | \mathcal{F}_t] \\
 E[X_u | \mathcal{F}_t] &= X_t + \mu(u - t),
 \end{aligned}$$

which violates a martingale condition. □

**Theorem 2.6 Detrended Brownian motion with drift is a continuous martingale**

Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, a detrended Brownian motion with drift defined as:

$$(X_{t \in [0, \infty)} - \mu t) \equiv (\mu t + \sigma B_{t \in [0, \infty)} - \mu t) \equiv (\sigma B_{t \in [0, \infty)}),$$

is a continuous martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

Proof

For  $\forall 0 \leq t \leq u < \infty$  :

$$\begin{aligned} E[X_u - \mu u | \mathcal{F}_t] &= E[(X_t - \mu t) + (\int_t^u dX_v - \mu \int_t^u dv) | \mathcal{F}_t] \\ E[X_u - \mu u | \mathcal{F}_t] &= E[(X_t - \mu t) | \mathcal{F}_t] + E[(\int_t^u dX_v - \mu \int_t^u dv) | \mathcal{F}_t] \\ E[X_u - \mu u | \mathcal{F}_t] &= X_t - \mu t + \mu(u - t) - \mu(u - t) \\ E[X_u - \mu u | \mathcal{F}_t] &= X_t - \mu t, \end{aligned}$$

which satisfies a martingale condition. □

**[2.3] A Continuous Martingale Property of Exponential Standard Brownian Motion**

**Theorem 2.7 Exponential of a standard Brownian motion is a continuous martingale**

Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion process defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, for any  $\theta \in \mathbb{R}$ , the exponential of a standard Brownian motion defined as:

$$Z_t = \exp(\theta B_t - \frac{1}{2} \theta^2 t), \tag{3}$$

is a continuous martingale with respect to the filtration  $\mathcal{F}_{t \in [0, \infty)}$  and the probability measure  $\mathbb{P}$ .

Proof

We first prove the often used proposition.

**Proposition 2.1** Note that if  $X \sim Normal(\mu t, \sigma^2 t)$ , then for any  $\theta \in \mathbb{R}$  :

$$E[\exp(\theta X)] = \exp\left(\theta\mu t + \frac{1}{2}\theta^2\sigma^2 t\right). \quad (4)$$

Proof

$$\begin{aligned} E[\exp(\theta X)] &= \int_{-\infty}^{\infty} \exp(\theta X) \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{(X - \mu t)^2}{2\sigma^2 t}\right\} dX \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{-\theta X 2\sigma^2 t + X^2 - 2X\mu t + \mu^2 t^2}{2\sigma^2 t}\right\} dX \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{X^2 - 2(\theta\sigma^2 t + \mu t)X + \mu^2 t^2}{2\sigma^2 t}\right\} dX \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{(X - (\theta\sigma^2 t + \mu t))^2 - (\theta\sigma^2 t + \mu t)^2 + \mu^2 t^2}{2\sigma^2 t}\right\} dX \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{(X - (\theta\sigma^2 t + \mu t))^2}{2\sigma^2 t}\right\} \exp\left\{\frac{(\theta\sigma^2 t + \mu t)^2 - \mu^2 t^2}{2\sigma^2 t}\right\} dX \\ &= \exp\left\{\frac{(\theta\sigma^2 t + \mu t)^2 - \mu^2 t^2}{2\sigma^2 t}\right\} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi\sigma^2 t}} \exp\left\{-\frac{(X - (\theta\sigma^2 t + \mu t))^2}{2\sigma^2 t}\right\} dX \\ &= \exp\left\{\frac{(\theta\sigma^2 t + \mu t)^2 - \mu^2 t^2}{2\sigma^2 t}\right\} = \exp\left\{\frac{\theta^2\sigma^4 t^2 + 2\theta\sigma^2 t\mu t}{2\sigma^2 t}\right\} \\ &= \exp\left\{\theta\mu t + \frac{1}{2}\theta^2\sigma^2 t\right\} \end{aligned}$$

□

Now we are ready to prove the Brownian exponential defined by the equation (3) is a martingale.

Firstly, the process  $(Z_{t \in [0, \infty)})$  is nonanticipating because a standard Brownian motion  $(B_{t \in [0, \infty)})$  is nonanticipating.

Secondly, it satisfies the finite mean condition, since  $E[Z_t] = 1 < \infty$  :

$$\begin{aligned} E[Z_t] &= E\left[\exp\left(\theta B_t - \frac{1}{2}\theta^2 t\right)\right] \\ E[Z_t] &= E\left[\exp(\theta B_t) \exp\left(-\frac{1}{2}\theta^2 t\right)\right] \\ E[Z_t] &= \exp\left(-\frac{1}{2}\theta^2 t\right) E\left[\exp(\theta B_t)\right], \end{aligned}$$

using the proposition 2.1,  $E[\exp(\theta B_t)] = \exp\left(\frac{1}{2}\theta^2 t\right)$  :

$$E[Z_t] = \exp(-\frac{1}{2}\theta^2 t) \exp(\frac{1}{2}\theta^2 t) = 1.$$

For  $\forall 0 \leq t \leq t+h < \infty$ , by the definition of  $Z_t$ :

$$E[Z_{t+h} | \mathcal{F}_t] = E[\exp\{\theta B_{t+h} - \frac{1}{2}\theta^2(t+h)\} | \mathcal{F}_t].$$

The trick is to multiply  $\exp(\theta B_t - \theta B_t) = e^0 = 1$  inside the expectation operator:

$$E[Z_{t+h} | \mathcal{F}_t] = E[\exp(\theta B_t - \theta B_t) \exp\{\theta B_{t+h} - \frac{1}{2}\theta^2(t+h)\} | \mathcal{F}_t]$$

$$E[Z_{t+h} | \mathcal{F}_t] = E[\exp(\theta B_t) \exp(-\theta B_t) \exp(\theta B_{t+h}) \exp(-\frac{1}{2}\theta^2 t) \exp(-\frac{1}{2}\theta^2 h) | \mathcal{F}_t]$$

$$E[Z_{t+h} | \mathcal{F}_t] = E[\exp(\theta B_t - \frac{1}{2}\theta^2 t) \exp\{\theta(B_{t+h} - B_t) - \frac{1}{2}\theta^2 h\} | \mathcal{F}_t].$$

Since Brownian increments are independent:

$$E[Z_{t+h} | \mathcal{F}_t] = E[\exp(\theta B_t - \frac{1}{2}\theta^2 t) | \mathcal{F}_t] E[\exp\{\theta(B_{t+h} - B_t) - \frac{1}{2}\theta^2 h\} | \mathcal{F}_t],$$

and since  $B_t$  is  $\mathcal{F}_t$ -adapted:

$$E[Z_{t+h} | \mathcal{F}_t] = \exp(\theta B_t - \frac{1}{2}\theta^2 t) E[\exp\{\theta(B_{t+h} - B_t) - \frac{1}{2}\theta^2 h\}].$$

By the definition of  $Z_t$ :

$$E[Z_{t+h} | \mathcal{F}_t] = Z_t E[\exp\{\theta(B_{t+h} - B_t)\} \exp(-\frac{1}{2}\theta^2 h)],$$

and since  $\exp(-\frac{1}{2}\theta^2 h)$  is a constant:

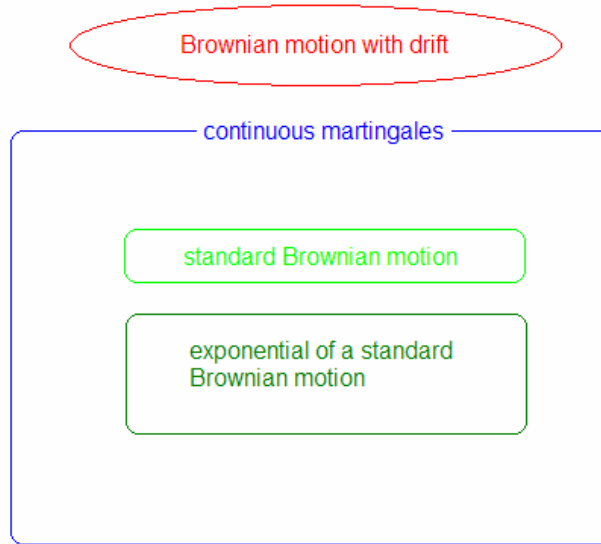
$$E[Z_{t+h} | \mathcal{F}_t] = Z_t \exp(-\frac{1}{2}\theta^2 h) E[\exp\{\theta(B_{t+h} - B_t)\}].$$

Use the proposition 2.1 because  $B_{t+h} - B_t \sim \text{Normal}(0, h)$ :

$$E[Z_{t+h} | \mathcal{F}_t] = Z_t \exp(-\frac{1}{2} \theta^2 h) \exp(\frac{1}{2} \theta^2 h)$$

$$E[Z_{t+h} | \mathcal{F}_t] = Z_t.$$

□



**Figure 2.1** Brownian motion as a subclass of continuous martingales

**[3] Brownian Motion as a Subclass of Gaussian Processes**

**Definition 3.1 Gaussian process** A stochastic process  $(X_{t \in [0, \infty)})$  on  $\mathbb{R}^d$  (i.e. this means that  $X_t$  is a  $d$ -dimensional vector) defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  is said to be a Gaussian process, if, for any increasing sequence of time  $0 \leq t_1 < t_2 < \dots < t_k < \infty$ , the law of any finite dimensional vector  $(X(t_1), X(t_2), \dots, X(t_k))$  of the process is multivariate normal.

Because all finite dimensional multivariate normal distributions are uniquely determined by their means and covariance function, Gaussian processes can be defined in an alternate way.

**Definition 3.2 Gaussian process** A stochastic process  $(X_{t \in [0, \infty)})$  on  $\mathbb{R}^d$  (i.e. this means that  $X_t$  is a  $d$ -dimensional vector) defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  is said to be a Gaussian process, if the law of the process  $(X_{t \in [0, \infty)})$  is uniquely determined by:

- (1) Means  $E[X_t]$ .

(2) Covariance functions  $Cov(X_t, X_u) = E[\{X_t - E(X_t)\}\{X_u - E(X_u)\}^T]$  for  $\forall 0 \leq t \neq u < \infty$ ,

where  $T$  is a transposition operator.

**Theorem 3.1 A standard Brownian motion** A standard Brownian motion  $(B_{t \in [0, \infty)})$  is a one dimensional Gaussian process with:

- (1) Zero mean  $E[B_t] = 0$ .
- (2) Covariance function  $Cov(B_t, B_u) = t \wedge u = \min\{t, u\}$ .

Proof

By the definition of a standard Brownian motion. For more details, we recommend Karlin and Taylor (1975) page 376-377.

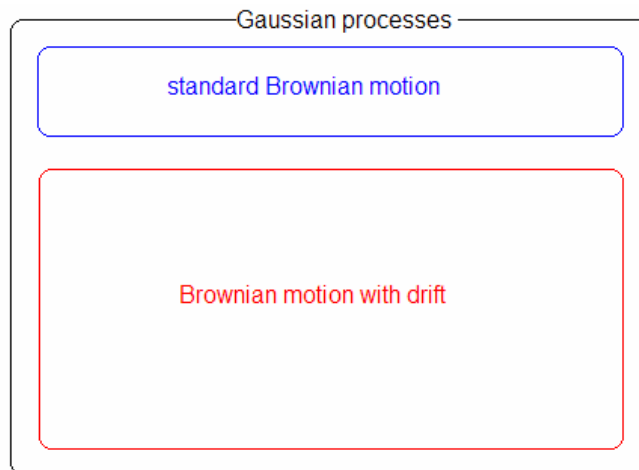
The converse of the theorem 3.1 is also true.

**Theorem 3.2 A standard Brownian motion** A real valued one dimensional Gaussian stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  is a standard Brownian motion with drift, if its mean and covariance function satisfy:

- (1)  $E[X_t] = 0$ .
- (2)  $Cov(X_t, X_u) = t \wedge u = \min\{t, u\}$

Proof

By the definition of a standard Brownian motion.



**Figure 3.1 Brownian motion as a subclass of Gaussian processes**

#### [4] Brownian Motion as a Subclass of Markov Processes

We first introduce the definitions and terminologies used in the study of Markov processes.

**Definition 4.1 Transition function** Consider a continuous time nonanticipating stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  which takes values in a measurable space  $(B, \mathcal{B})$  (i.e.  $B \in \mathcal{B}(\mathbb{R})$ ).  $(B, \mathcal{B})$  is called a state space of the process and the process is said to be  $B$ -valued. Consider an increasing sequence of time  $0 \leq t \leq u \leq v < \infty$ . A real valued transition function  $\mathbb{P}_{t,v}(x, B)$  with  $x \in \mathbb{R}$  and  $B \in \mathcal{B}(\mathbb{R})$  is a mapping which satisfies the following conditions:

- (1)  $\mathbb{P}_{t,v}(x, B)$  is a probability measure which maps every fixed  $x$  into  $B$ .
- (2)  $\mathbb{P}_{t,v}(x, B)$  is  $\mathcal{B}$ -measurable for every  $B \in \mathcal{B}(\mathbb{R})$ .
- (3)  $\mathbb{P}_{t,t}(x, B) = \delta(B)$ .
- (4)  $\mathbb{P}_{t,v}(x, B) = \int_{\mathbb{R}} \mathbb{P}_{t,u}(x, dy) \mathbb{P}_{u,v}(y, B)$ .

The condition (4) is called the Chapman-Kolmogorov identity. Chapman-Kolmogorov identity means that the transition probability  $\mathbb{P}_{t,v}(x, B)$  of moving from a state  $x$  at time  $t$  to a state  $B$  at time  $v$  can be calculated as a sum (i.e. integral) of the product of the transition probabilities via an intermediate time  $t \leq u \leq v$ , i.e.  $\mathbb{P}_{t,u}(x, dy)$  and  $\mathbb{P}_{u,v}(y, B)$ . In the general cases, transition functions are dependent on the states and time.

**Definition 4.2 Time homogeneous (temporary homogeneous or stationary) transition function** Consider an increasing sequence of time  $0 \leq t \leq u \leq v < \infty$ . A real valued transition function  $\mathbb{P}_{t,v}(x, B)$  with  $x \in \mathbb{R}$  and  $B \in \mathcal{B}(\mathbb{R})$  is said to be time homogeneous if it satisfies:

$$\mathbb{P}_{t,v}(x, B) = \mathbb{P}_{0, v-t}(x, B) = \mathbb{P}_{v-t}(x, B),$$

which indicates that the transition function  $\mathbb{P}_{t,v}(x, B)$  of moving from a state  $x$  at time  $t$  to a state  $B$  at time  $v$  is equivalent to the transition function  $\mathbb{P}_{0, v-t}(x, B)$  of moving from a state  $x$  at time 0 to a state  $B$  at time  $v-t$ . In other words, the transition function is independent of the time  $t$  and depends only on the interval of time  $v-t$ .

**Definition 4.3 Chapman-Kolmogorov identity for the time homogeneous transition function** Consider an increasing sequence of time  $0 \leq t \leq u < \infty$ . Chapman-Kolmogorov identity for the time homogeneous transition function is:

$$\int_{\mathbb{R}} \mathbb{P}_{0,t}(x, dy) \mathbb{P}_{0,u}(y, B) = \int_{\mathbb{R}} \mathbb{P}_t(x, dy) \mathbb{P}_u(y, B) = \mathbb{P}_{t+u}(x, B).$$

**Definition 4.4 Markov Processes (less formal)** Consider a continuous time nonanticipating stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Then, the process  $(X_{t \in [0, \infty)})$  is said to be a Markov process if it satisfies, for every increasing sequence of time  $0 < t_1 \leq t_2 \leq \dots \leq t_n \leq t \leq u < \infty$ :

$$\mathbb{P}(X_u | \mathcal{F}_t) = \mathbb{P}(X_u | X_0, X_{t_1}, X_{t_2}, \dots, X_{t_n}, X_t) = \mathbb{P}(X_u | X_t),$$

Informally, Markov property means that the probability of a random variable  $X_u$  at time  $u \geq t$  (tomorrow) conditional on the entire history of the stochastic process  $\mathcal{F}_{[0, t]} \equiv X_{[0, t]}$  is equal to the probability of a random variable  $X_u$  at time  $u \geq t$  (tomorrow) conditional only on the value of a random variable at time  $t$  (today). In other words, the history (sample path) of the stochastic process  $\mathcal{F}_{[0, t]}$  is of no importance in that the way this stochastic process evolved or the dynamics does not mean a thing in terms of the conditional probability of the process. This property is sometimes called a memoryless property.

**Definition 4.5 Markov Processes (formal)** Consider a continuous time nonanticipating stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  which takes values in a measurable space  $(B, \mathcal{B})$ .  $(B, \mathcal{B})$  is called a state space of the process and the process is said to be  $B$ -valued. Then, the process  $(X_{t \in [0, \infty)})$  is said to be a Markov process if it satisfies, for an increasing sequence of time  $0 \leq t \leq u \leq v < \infty$   $0 < t \leq u < \infty$ :

$$E[X_v | \mathcal{F}_t] = E[X_v | X_t],$$

with the transition function:

$$\mathbb{P}_{t, v}(x, B) = \int_{\mathbb{R}} \mathbb{P}_{t, u}(x, dy) \mathbb{P}_{u, v}(y, B).$$

Now we are ready to characterize a standard Brownian motion as a subclass of Markov processes.

**Theorem 4.1 A standard Brownian motion process** A standard Brownian motion  $(B_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  satisfies the followings:

(1) It is a time homogeneous Markov process. In other words, for any bounded Borel function  $f : \mathbb{R} \rightarrow \mathbb{R}$  and for  $\forall 0 \leq t \leq u < \infty$ :

$$E[f(B_u) | \mathcal{F}_t] = \mathbb{P}_{0, u-t} f(B_t) = \mathbb{P}_{u-t} f(B_t).$$



(2) Its transition function  $\mathbb{P}_{u-t} \equiv \mathbb{P}_h$  is given by:

$$\mathbb{P}_h(x, y) = \frac{1}{\sqrt{2\pi t}} \exp\left\{-\frac{(x-y)^2}{2t}\right\}.$$

$$(3) \mathbb{P}_h f(x) = \begin{cases} f(x) & \text{if } t=0 \\ \int_{-\infty}^{\infty} \mathbb{P}_h(x, y) f(y) dy & \text{if } t>0 \end{cases}.$$

Proof

Markov property is a result of independent increments property of Brownian motion. Let  $(B_{t \in [0, \infty)})$  be a standard Brownian motion defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$ . Consider an increasing sequence of time  $0 < t_1 < t_2 < \dots < t_n < t < u < \infty$  where  $t$  is the present. As a result of independent increments condition:

$$\begin{aligned} & \mathbb{P}(X_u - X_t \mid X_{t_1} - X_0, X_{t_2} - X_{t_1}, \dots, X_t - X_{t_n}) \\ &= \frac{\mathbb{P}(X_u - X_t \cap X_{t_1} - X_0, X_{t_2} - X_{t_1}, \dots, X_t - X_{t_n})}{\mathbb{P}(X_{t_1} - X_0, X_{t_2} - X_{t_1}, \dots, X_t - X_{t_n})} \\ &= \frac{\mathbb{P}(X_u - X_t) \mathbb{P}(X_{t_1} - X_0, X_{t_2} - X_{t_1}, \dots, X_t - X_{t_n})}{\mathbb{P}(X_{t_1} - X_0, X_{t_2} - X_{t_1}, \dots, X_t - X_{t_n})} \\ &= \mathbb{P}(X_u - X_t), \end{aligned}$$

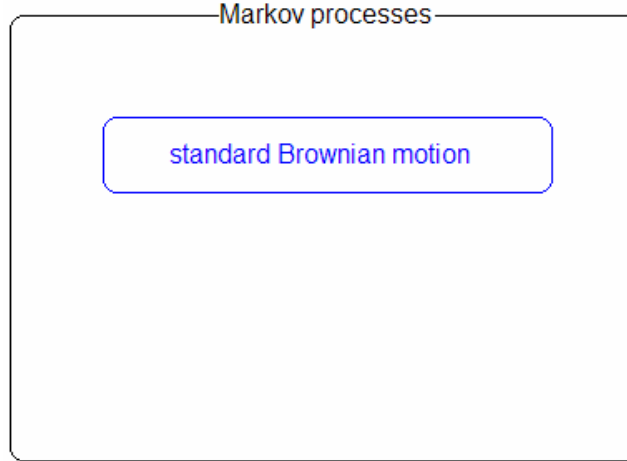
which means that there is no correlation (probabilistic dependence structure) on the increments among the past, the present, and the future.

Using the simple relationship  $X_u \equiv (X_u - X_t) + X_t$  for an increasing sequence of time  $0 < t_1 < t_2 < \dots < t_n < t < u < \infty$ :

$$\begin{aligned} \mathbb{P}(X_u \mid X_0, X_{t_1}, X_{t_2}, \dots, X_{t_n}, X_t) &= \mathbb{P}((X_u - X_t) + X_t \mid X_0, X_{t_1}, X_{t_2}, \dots, X_{t_n}, X_t) \\ &= \mathbb{P}(X_u \mid X_t), \end{aligned}$$

which holds because an increment  $(X_u - X_t)$  is independent of  $X_t$  by definition and the value of  $X_t$  depends on its realization  $X_t(\omega)$ .

□



**Figure 4.1** Brownian motion as a subclass of Markov processes

**[5] Brownian Motion as a Subclass of Itô Diffusion Processes**

**Definition 5.1 Itô diffusion processes** An Itô diffusion process is a real valued stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  whose dynamics (or motion) is governed by a stochastic differential equation of the form:

$$dX_t = b(t, X_t)dt + \sigma(t, X_t)dB_t,$$

where  $b(t, X_t) \in \mathbb{R}$  is called a drift and  $\sigma(t, X_t)$  which is a nonnegative real valued constant is called a diffusion parameter. In the general case,  $b(t, X_t)$  and  $\sigma(t, X_t)$  are functions of both time and space. As usual,  $B_t$  stands for a standard Brownian motion. The solution of the above stochastic differential equation is given by:

$$X_t = X_0 + \int_0^t b(s, X_s)ds + \int_0^t \sigma(s, X_s)dB_s.$$

**Definition 5.2 Time homogeneous (temporary homogeneous or stationary) Itô diffusion processes** A time homogeneous Itô diffusion process is a real valued stochastic process  $(X_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  whose dynamics (or motion) is governed by a stochastic differential equation of the form:

$$dX_t = b(X_t)dt + \sigma(X_t)dB_t,$$

where a drift  $b(X_t) \in \mathbb{R}$  and a diffusion parameter  $\sigma(X_t) \geq 0$  are independent of the time  $t$  and depend only on the space.

**Theorem 5.1 Time homogeneous Itô diffusion processes are a subclass of time homogeneous Markov processes** A time homogeneous Itô diffusion process  $(X_{t \in [0, \infty)})$

defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  is a time homogeneous Markov process. In other words, for any bounded Borel function  $f : \mathbb{R} \rightarrow \mathbb{R}$  and for  $\forall 0 \leq t \leq u < \infty$  :

$$E[f(B_u) | \mathcal{F}_t] = \mathbb{P}_{0, u-t} f(B_t) = \mathbb{P}_{u-t} f(B_t),$$

where  $\mathbb{P}_{u-t} \equiv \mathbb{P}_h$  is a time homogeneous transition function given by:

$$\mathbb{P}_h f(x) = \begin{cases} f(x) & \text{if } t = 0 \\ \int_{-\infty}^{\infty} \mathbb{P}_h(x, y) f(y) dy & \text{if } t > 0 \end{cases}.$$

Proof

Consult Oksendal (2003) pages 115-116.

**Theorem 5.2 A standard Brownian motion** A standard Brownian motion  $(B_{t \in [0, \infty)})$  defined on a filtered probability space  $(\Omega, \mathcal{F}_{t \in [0, \infty)}, \mathbb{P})$  is a time homogeneous Itô diffusion process  $(X_{t \in [0, \infty)})$  (a time homogeneous Markov process) whose dynamics (or motion) is governed by a stochastic differential equation of the form:

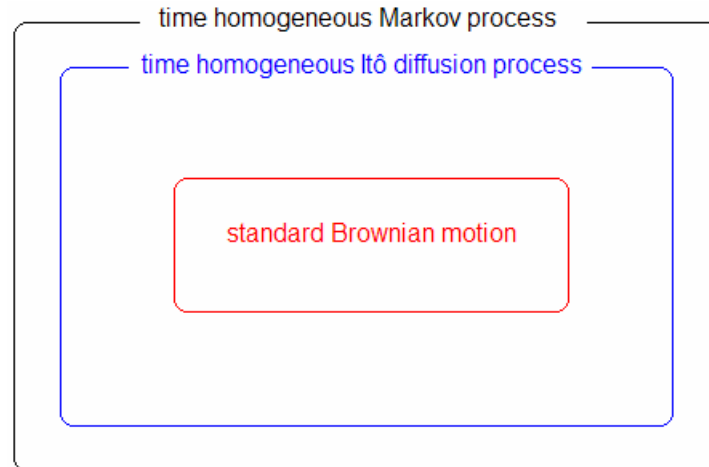
$$dX_t = dB_t.$$

In other words, a standard Brownian motion  $(B_{t \in [0, \infty)})$  is a time homogeneous Itô diffusion process with the zero drift  $b(t, X_t) = 0$  and the unit diffusion parameter  $\sigma(t, X_t) = 1$ .

Proof

By the definition of a standard Brownian motion.

For more details about Itô diffusion processes, consult an excellent book Oksendal (2003) chapter 3 and 7.



**Figure 5.1 Brownian motion as a subclass of Markov processes**

### References

Karlin, S., and Taylor, H., 1975, A First Course in Stochastic Processes, Academic Press.

Karatzas, Ioannis., and Shreve, S. E., 1991, Brownian Motion and Stochastic Calculus, Springer-Verlag.

Oksendal, B., 2003, Stochastic Differential Equations: An Introduction with Applications, Springer.

Rogers, L.C.G., and Williams, D., 2000, Diffusions, Markov Processes and Martingales, Cambridge University Press.

Sato, K., 1999, Lévy process and Infinitely Divisible Distributions, Cambridge University Press.